

The impact of human mobility networks on the global spread of COVID-19

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Human mobility networks are crucial for a better understanding and controlling the spread of epidemics. Here, we study the impact of human mobility networks on the COVID-19 onset in 203 different countries. We use exponential random graph models to perform an analysis of the country-to-country global spread of COVID-19. We find that most countries had similar levels of virus spreading, with only a few acting as the main global transmitters. Our evidence suggests that migration and tourism inflows increase the probability of COVID-19 case importations while controlling for contiguity, continent co-location and sharing a language. Moreover, we find that air flights were the dominant mode of transportation while male and returning travellers were the main carriers. In conclusion, a mix of mobility and geography factors predicts the COVID-19 global transmission from one country to another. These findings have implications for non-pharmaceutical public health interventions and the management of transborder human circulation.

Keywords: COVID-19 case importations; human mobility networks; incoming migration; inbound tourism; complex networks.

1. Introduction

The ongoing pandemic of coronavirus disease (COVID-19), caused by the severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2), originated in Wuhan, China. The SARS-CoV-2, a coronavirus that infects humans, spread, as of September 2020, in 216 countries and territories [1]. COVID-19 has proved to be a highly contagious pathogenic viral infection, globally amassing more than 30 million confirmed cases and giving rise to approximately 1 million deaths [1], since its detection in December 2019. Subsequent non-pharmaceutical interventions (NPIs) have been implemented by both authorities in China and worldwide to decrease the number of new daily infections and deaths, as well as to control the demand for healthcare and avoid the overwhelming of national medical systems. Essentially, many of these interventions have consistently restricted human mobility. Either with the purpose to reduce local transmissions,

for example, the case of physical distancing—closure of schools and universities, of workplaces, isolation, lockdown, ban of gatherings, etc. Or to decline regional, national and international transmissions, for example, the case of international and domestic travelling restrictions—travel bans, quarantine, border closure, etc. The quantitative investigation of the NPIs' efficiency and the effectiveness in response to COVID-19 is currently underway [2, 3]. On one hand, several studies argue that a predetermined change in the human mobility patterns is expected to delay [4, 5], mitigate [6] or reduce [7] the COVID-19 spread. Recent developments even suggest that individual-level travel history and demographic characteristics can be used to predict the spread and size of the COVID-19 in China [8]. On the other hand, some claim it is unclear whether the restricting movement has been a proportional response to the COVID-19 pandemic [9] and that further research on the topic is needed.

The importance of human mobility patterns in understanding and controlling the spread of biological viruses has already been documented [10–12]. Even more, it has been claimed that the diversity of human circulation can be reduced to simple stable patterns which, in turn, become relevant for phenomena such as epidemic prevention or response [13]. Our study starts from the already reported evidence showing the contribution of carriers (returned travellers and visitors) to the international spread of COVID-19, during its early stage [14–17]. A plethora of case studies illustrates that country-level COVID-19 dispersion was ignited by case importations [18–26], and international migration [27]. However, little is known on the country-to-country global spread of COVID-19, and a description is missing for the moment. Available studies have tended to focus on whether air travel volume data could predict COVID-19 outbreaks [28–30]. Additionally, this line of work has not assessed the effect of human mobility *per se*, in explaining the COVID-19 cases onset in each country worldwide.

In this article, the first objective is to cover the current gap in the literature by describing the early COVID-19 global transmission from one country to another. The second objective is to contribute to the understanding of the COVID-19 spread by examining the associated role of human mobility networks. To these objectives, our attention is not focused on the medical and biological features of COVID-19. But, we draw our investigation stemming from the theoretical reasoning that the spread of epidemics among humans is inherently a social network problem. Contact tracing, social interaction patterns and mobility fluxes feed epidemic forecast and modelling [10]. Consequently, we look at two forms of global human mobility: inflows of migration and tourism. We hypothesize that these factors (incoming migration—hypothesis 1) and (inbound tourism—hypothesis 2) have a positive impact upon the initial global spread of COVID-19 (the dependent variable). To test these hypotheses, we employ a socio-centric network analysis research design and fit exponential random graph models (ERGMs) [31]. In our analysis, the dependent variable is built after having collected a unique dataset on the first cases of COVID-19 officially reported by national authorities. Both the predictors of interest and dependent variable are input into the models as origin–destination matrices. In our statistical modelling, we also control for possible confounding variables, such as geographical (contiguity and sharing the same continent) and cultural proximity (sharing a common language). These variables, included as symmetric binary matrices, may affect both the migration [32, 33] and tourism [34–36] inflows as well as the initial COVID-19 spread [27]. All variables included in the ERGM models are real-world country-level network data.

2. Results

We start by reporting descriptive statistics on the individual COVID-19 case onset at a country-level. Throughout the 219 countries and territories, we observe 323 COVID-19 cases (some countries reported more than one patient as the first officially confirmed case), with an average age of 45.9 (95% CI 44.7–47.1, SD=16.5, SE=1.2, missing = 132). Among these, 94 (29.1%) are females, 156 (48.3%) are males,

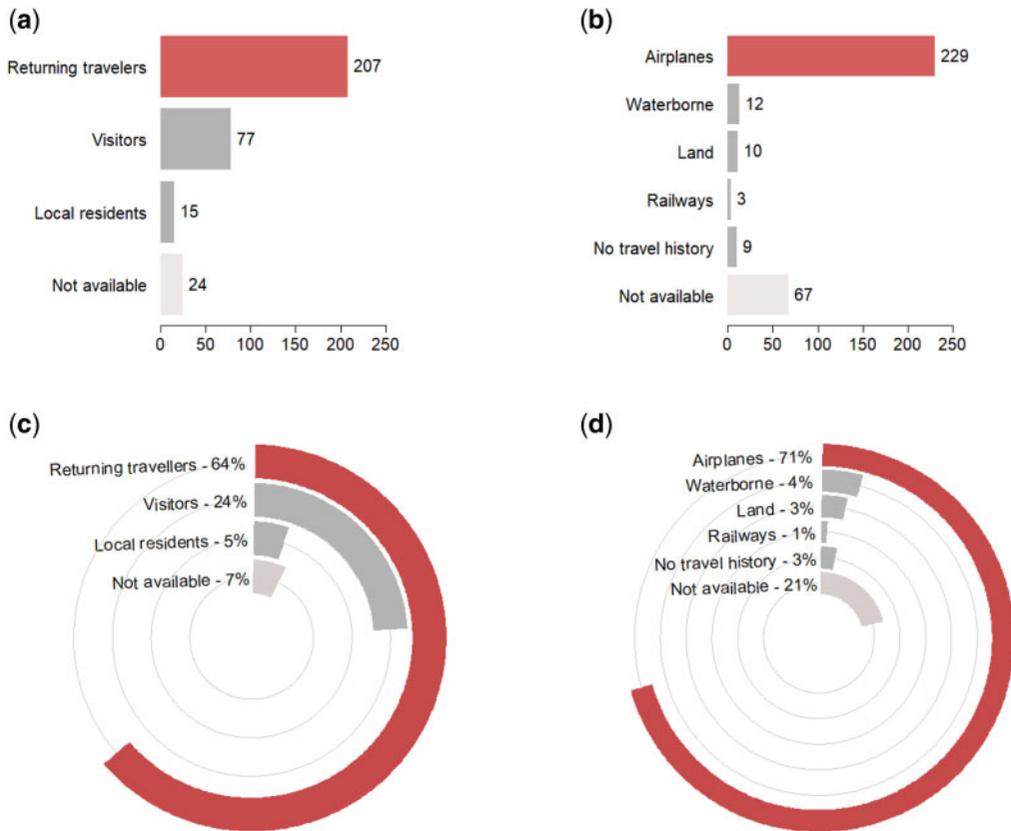


FIG. 1. Type of travellers and modes of transportation for the first 323 individual COVID-19 cases. Images (a) and (b) are counts, while (c) and (d) are percentages.

while for 73 (22.6%) patients there is no available evidence on the sex variable. In a 3-month time, COVID-19 spread in almost all the countries and territories of the globe. Specifically, approximately 67% of all first cases were officially announced during March 2020. Also, 95% of the countries and territories publicly confirmed their first cases between 1 January 2020 and 31 March 2020. Specifically, 12% in January, 18% in February and 65% in March. Put it differently, in 3 months, COVID-19 spread in 208 out of the 219 documented countries and territories.

Figure 1 displays the distribution of the first individual COVID-19 cases by type of travellers and by modes of transportation (where applicable). As shown in Fig. 1(a and c), of all the 323 documented cases, 64% ($n = 207$) are returning travellers, 24% ($n = 77$) visitors, 5% ($n = 15$) local residents and 7% ($n = 24$) are missing data (i.e. data not available in the country-level official reports). In terms of transportation modes (Fig. 1(b and d)), the airplanes were used by 71% ($n = 229$) of the individual cases, waterborne transport was used by 4% ($n = 12$), land transport (bus or car) was used by 3% ($n = 10$) and railways by 1% ($n = 3$). For 21% ($n = 67$) of the patients, the means of transportation was not public information (i.e. data not available). In the case of 3% ($n = 9$) of the cases, patients reported not having a recent travel history during the past 20 days. It is noteworthy that the percentage exceeds 100% as, in

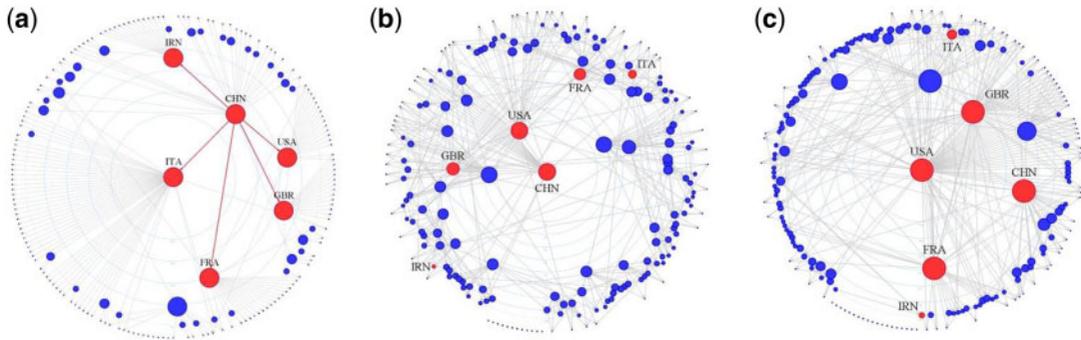


FIG. 2. Centrality layouts for visualizing the patterns of (a) COVID-19, (b) incoming migration and (c) inbound tourism ties among countries. Node size is proportional to out-degree centrality. Countries located more centrally in the pictures have higher centrality scores. Core-countries are marked with red. In (a), ties among core-countries are marked in red.

some cases ($n = 6$), multiple means of transportation were used, for example, railways and bus, airplane and bus or car, maritime and airplane.

Using out-degree centrality layouts, we illustrate three binary directed graphs: *the COVID-19 network* (203 nodes, 213 ties, density (D) = 0.005, out-degree centralization (*out-deg C*) = 0.249, in-degree centralization (*in-deg C*) = 0.015) (Fig. 2a), *the incoming migration network* (203 nodes, 571 ties, $D = 0.014$, *out-deg C* = 0.001, *in-deg C* = 0.105) (Fig. 2b) and *the inbound tourism network* (203 nodes, 519 ties, $D = 0.013$, *out-deg C* = 0.002, *in-deg C* = 0.321) (Fig. 2c). In Fig. 2, variations in country-level out-degree centrality scores (computed as percentage) is marked by both layout location (states with higher out-degrees are more central) and node size (the larger the node, the higher the centrality score). The centralization of the incoming migration network (both in-degree and out-degree) is influenced by the restrictions on the tie eliciting (only the first top three migration source countries were accounted in the analysis).

Fitting a core/periphery model [37, 38], we found that *China (CHN)*, the *United States of America (USA)*, *France (FRA)*, *Italy (ITA)*, *Great Britain (GRB)* and *Iran (IRN)* represent the core of the COVID-19 network. The core/periphery model has a fit (correlation) of 0.2, for an expected density of *core/periphery ties* and *periphery/core ties* of 0.5 (at five random starts and 100 iterations). In Fig. 2, we mark the six core-countries with red in all the three networks, for rapid identification of their centrality placement. All three networks fit power-law distributions [39]. Particularly, this is indicated by Kolmogorov–Smirnov (KS) statistic tests, that is, the COVID-19 network ($KS = 0.07$, $p = 0.99$), the incoming migration network ($KS = 0.05$, $p = 0.99$), the inbound tourism network ($KS = 0.04$, $p = 0.99$).

Figure 3(a–c) displays the (*common-*)*language-network* (203 nodes, 3194 edges, $D = 0.156$, *deg C* = 0.308), the *contiguity-network* (203 nodes, 301 edges, $D = 0.015$, *deg C* = 0.065) and the (*location-*)*same-continent-network* (203 nodes, 4756 edges, $D = 0.232$, *deg C* = 0.031). The language-network has 40 components (the main component, *MC*, of 152 nodes, four components of triads, four dyads and 31 isolates), while the contiguity-network has 52 components (an *MC* of 128 nodes, another component of 23 nodes, 2 dyads and 48 isolates). The same-continent-network is made up of five components, corresponding to the geographical regions designed by the United Nations. Specifically, the countries and territories included in the analysis are spread across the regions in the following way: Africa—54 nodes, America—50 nodes, Asia—49, Europe—43 and Oceania—7. The groupings of countries and territories by regions can be scrutinized in Fig. 3c.

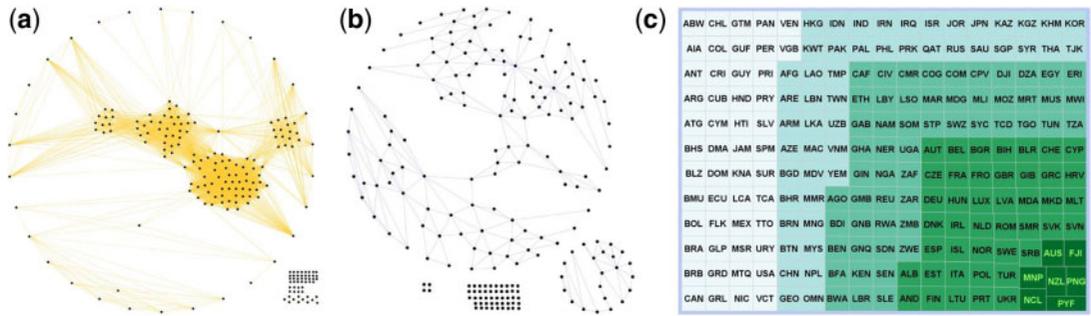


FIG. 3. Visualizations of (a) the common language network, (b) the contiguity network and (c) the location on the same continent network. In (c), countries sharing the same geographical region are marked by colour (we employed United Nations’ region classification: America, Asia, Africa, Europe and Oceania). Also, the states included in the network are available by their ISO-alpha three codes (three-letter abbreviation).

TABLE 1 QAP Pearson’s correlations

Network variables	1	2	3	4	5
1. COVID-19 network					
2. Migration	0.049 ($p = 0.0002$) SD = 0.0048				
3. Tourism	0.013 ($p = 0.0144$) SD = 0.0046	0.355 ($p = 0.0002$) SD = 0.0060			
4. Contiguity	0.065 ($p = 0.0002$) SD = 0.0053	0.406 ($p = 0.0002$) SD = 0.0056	0.340 ($p = 0.0002$) SD = 0.0058		
5. Common language	0.038 ($p = 0.0002$) SD = 0.0095	0.102 ($p = 0.0002$) SD = 0.0065	0.101 ($p = 0.0002$) SD = 0.0097	0.097 ($p = 0.0002$) SD = 0.0081	
6. Same continent	0.032 ($p = 0.0002$) SD = 0.0052	0.156 ($p = 0.0002$) SD = 0.0056	0.128 ($p = 0.0002$) SD = 0.0055	0.206 ($p = 0.0002$) SD = 0.0070	0.165 ($p = 0.0002$) SD = 0.0077

Correlations were performed based on 5000 permutations.

Table 1 reports the quadratic assignment procedure (QAP) Pearson’s correlation coefficients computed for the squared adjacency matrices previously illustrated as networks in Figs. 2 and 3 [38]. All coefficients are positive and statistically significant. The pair-wise correlation measurements indicate rather weak relationships between the network variables.

Table 2 presents the most important results of our article: the estimates for the structural and relational attribute effects included in the ERGM models (see Section 4 for the definition of the effects). These models differ in the following way. Model 1 examines purely structural factors (*edges*, *activity spread*, *indegree* of 2, 1 and 0, *outdegree* of 2, 1 and 0, *twopaths*, and *isolates*). Model 2 examines the *incoming migration* and *inbound migration* effects (the predictors or the relational attribute effects of interest), controlling for structural effects. Model 3 (the full model) examines the predictors of interests, while

TABLE 2 ERGM estimates of structural and relational attribute effects on the presence of a COVID-19 tie between a sending and a receiving country

	Model 1			Model 2			Model 3		
	Est.	<i>p</i> -value	SE	Est.	<i>p</i> -value	SE	Est.	<i>p</i> -value	SE
Structural effects									
Edges	-1.70	0.0515	(0.87)	-2.09	0.0398	(1.02)	-2.36	0.0159	(0.98)
Activity spread [†]	-4.84	0.0001	(0.38)	-3.96	0.0001	(0.39)	-4.30	0.0001	(0.40)
Indegree (2)	0.28	0.8229	(1.26)	0.32	0.8211	(1.41)	0.31	0.8172	(1.35)
Indegree (1)	2.23	0.2768	(2.06)	2.33	0.3224	(2.35)	2.27	0.3141	(2.25)
Indegree (0)	-0.44	0.8841	(3.01)	-0.27	0.9365	(3.43)	-0.40	0.9041	(3.28)
Outdegree (2)	-1.26	0.0713	(0.70)	-0.89	0.2209	(0.72)	-1.06	0.1412	(0.72)
Outdegree (1)	-1.30	0.1039	(0.80)	-0.55	0.5050	(0.82)	-0.92	0.2648	(0.82)
Outdegree (0)	-0.59	0.5761	(1.05)	0.66	0.5442	(1.08)	0.03	0.9757	(1.09)
Twopath	-0.05	0.2356	(0.04)	-0.05	0.2621	(0.04)	-0.05	0.2335	(0.04)
Isolates	0.10	0.9076	(0.87)	0.10	0.9040	(0.84)	0.09	0.9107	(0.83)
Relational attribute effects									
Incoming migration				1.36	0.0001	(0.22)	0.97	0.0001	(0.23)
Inbound tourism				0.93	0.0001	(0.18)	0.77	0.0001	(0.16)
Contiguity							0.75	0.0057	(0.27)
Common language							0.73	0.0001	(0.14)
Same continent							0.37	0.0220	(0.16)
AIC	1631			1507			1476		
BIC	1717			1611			1605		
Log Likelihood	-805			-742			-723		

[†]Decay parameter fixed at 2, for activity spread (two-degree).

controlling both for the structural effects and for the other three relational attribute effects (*contiguity*, *common language* and *same continent*).

We notice, in Table 2, that *incoming migration* and *inbound tourism* have a statistically significant positive impact ($p = 0.0001$) in the process of COVID-19 tie formation (Models 2 and 3). Specifically, COVID-19 case importations are more likely from states that are important sources of migration and tourism. Model 2 illustrates an increase of 3.9 (Exp(1.36), $Est = 1.36$, $p = 0.0001$, $SE = 0.22$) in the odds of receiving a COVID-19 tie from a country that is a top source of migration, and of 2.5 ($Est = 0.93$, $p = 0.0001$, $SE = 0.18$), from a country that is a top source of tourism. The probabilities of case importations from top sending migration and tourism countries are 0.80 ($\exp(1.36)/(1+\exp(1.36))$), and 0.72, respectively. Model 3 shows that the effects of incoming migration and inbound tourism remain statistically significant positive ($p = 0.0001$) while controlling for geographical (*contiguity* and *same continent*) and cultural (*common language*) proximities. Precisely, the corresponding probabilities of a COVID-19 case importation are 0.73 ($Est = 0.97$, $p = 0.0001$, $SE = 0.23$) and 0.68 ($Est = 0.77$, $p = 0.0001$, $SE = 0.16$). We also observe that geographical (*contiguity* and *same continent*) and cultural (*common language*) proximities are statistically significant predictors ($p = 0.0057$, $p = 0.0220$ and $p = 0.0001$, respectively) when accounting for COVID-19 ties between countries. Importing COVID-19 cases from neighbouring countries (*contiguity*) has a 0.68 probability ($Est = 0.75$, $p = 0.0057$, $SE = 0.27$), while COVID-19 contagion on the *same continent*, a probability of 0.59 ($Est = 0.37$, $p = 0.0220$, $SE = 0.16$). *Common language* gives an increase of 2.1 ($Est = 0.73$, $p = 0.0001$, $SE = 0.14$) in the odds of receiving a COVID-19 tie (the corresponding tie probability is 0.67).

Referring to the structural effects, the tendency of forming ties in the network (*edges*) has statistically significant negative estimates in Models 2 and 3 ($Est = -2.09$, $p = 0.0398$, $SE = 1.02$, and $Est = -2.36$,

95% of the countries and territories that officially confirmed cases. A global diffusion of this magnitude may indicate that national states lacked an efficient coordination in controlling the country-to-country virus spread. A core group of countries seems to have involuntarily acted as hubs in the global virus transmission, confirming one of the well-documented features of complex networks, the preferential attachment [50]. The out-degree centrality of these countries in other global structures, such as migration and tourism networks, may have facilitated their acting as main global transmitters.

In line with our research objectives stated in the Section 1, we described how COVID-19 spread from one country to another and demonstrated the statistically significant effect of migration and tourism inflows ($p = 0.0001$). One possible implication refers to the utility that models predicting population movements [13, 51] have in preparing responses to future potential different pandemics. Or even possible next waves of COVID-19. Modelling global human mobility may be an important input in increasing the international coordination of national authorities for handling cross-border circulation. Additionally, we believe our study may be a valuable contribution to developing global strategies for the early prevention of pandemics. For instance, coordination of information exchange between neighbouring countries as well as between states that share cultural and migration corridors may be of vital importance for case importation detection. National early warning systems should be internationally synchronized and integrate customized real-time information about global mobility flows as well as about countries that act as global transmitters. Furthermore, international coordination of information streams on people's movement patterns may be a driver for increasing the consistency level of the governmental preventive measures. It may be the case that the existing uncoordinated variations in the responses of domestic authorities to the COVID-19 to have facilitated the global virus spread. Lastly, our work may be particularly useful for researchers aiming to compare COVID-19 to other biological viruses' non-medical spreading dynamics, such as severe acute respiratory syndrome (SARS) for example [48]. On top of that, we hope our results may be beneficial for policymakers and national authorities in their efforts of managing human mobility across borders, economic difficulties and migration-related issues. Rapid action is needed as the COVID-19 pandemic has heavily hit the tourism industry and the air-flight sector. And, perhaps, even more important, the global spread of COVID-19 has consistently negatively impacted upon migration, exacerbating social vulnerabilities and polarities, increasing stigmatization and exclusion and aggravating migrants' health vulnerabilities.

We are aware that our study may have at least two limitations. The first refers to the official country-level identification of the first COVID-19 cases. Our raw data come from national authorities' public announcements. Theoretically, that does not eliminate the possibility of COVID-19 existing in a country prior to being officially reported. The second concerns the network variables included in our statistical modelling as control variables. The measurement of geographical and cultural proximities was limited to contiguity, co-location on the same continent and language sharing. This limit is due to the fact that, generally, similar proxy variables at a global level are exceptionally scarce. Despite these limitations, it is noteworthy that the estimates of our statistical models hold for any global COVID-19 network that displays similar structural features.

To the best of our knowledge, this is the first study to document the global spread of COVID-19 outbreak and to assess the effects of human mobility networks. The evidence from this study shows that the global spread of the COVID-19 outbreak was not random but patterned by the incoming migration and inbound tourism, with geographical and cultural proximities being controlled for. Further research on this topic is needed to increase our understanding of how nowadays global human mobility routes impact upon virus spreading.

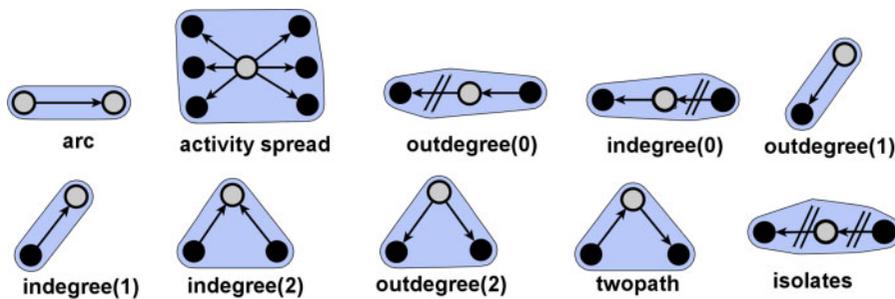


FIG. 5. Structural configurations included as effects in the ERGMs.

Triads similar to the one displayed as Fig. 4b were used to build the (incoming/ inflow) migration network and the (inbound) tourism network. Particularly, the top three migration sending countries and the top three tourism sending countries were identified for each of the states included in the COVID-19 network. We used the most recent available data on the global migration flows (2017), from the World Bank [52] and on the global tourism flows (2016), from World Tourism Organization [53].

In the *contiguity network*, two countries are connected by an edge if they share the same border. In the *common language network*, two states are linked if at least 20% of the population in one country speak the language of the other country. In the *same continent network*, two states are joined together if located on the same continent (we employed the geo-scheme advanced by the United Nations [54]). We used undirected dyads, similar to the one exhibited in Fig. 4c, to build these networks. The information on the geographical variables is publicly available from Centre D'Études Prospective Et D'Informations Internationales [55].

We used a suite of software tools for network descriptive statistics and visualization. Precisely, we run the *core/periphery* algorithm available with UCINET 6.0 [37, 38] for detecting the core-countries in the COVID-19 network. Also, we performed the *Kolmogorov–Smirnov* test for power-law distribution fit, using the algorithm available with the NetIndices R-package [56]. For network visualizations, we employed the circular and treemap node layouts available in visone [57]. The dataset files used in this article as well as the code are available for replication [58].

Initially, we performed quadratic assignment procedure (QAP) correlations (with 5000 permutations) to infer association among network variables, at a dyadic level. QAP correlations were useful to identify multi-collinearity, but not to explain patterns of ties. Consequently, we modelled the COVID-19 network using exponential random graph models (ERGMs) [31]. Generally, ERGMs are statistical models for understanding how and why network ties arise in a specific network [41]. In this article, we were interested in examining the generative processes giving rise to the structures or tie patterns in the observed network of COVID-19 countries. In our ERGMs, we estimated two classes of predictors: structural effects and relational attribute effects. The coefficients estimated for each predictor define the probability of each edge in the COVID-19 network and the probability of the entire network [59, 60].

The *structural effects* refer to specific tie patterns or link configurations within the network [59, 60]. In Fig. 5, we visually display the structural effects introduced in our ERGMs. These images are built on similar visualizations available in the literature [41]. *Edges (arcs)* refer to the overall tendency for COVID-19 ties to form in the network. Generally, ERGMs include this term by default. *Activity spread* term controls for the COVID-19 network's tendency to have active nodes (countries that send many arrows).

The *out-degree (0)* or *sinks* parameter marks nodes with zero out-degree but positive in-degree. The *in-degree (0)* or *sources* parameter marks nodes with zero in-degree but positive out-degrees. The *isolates* term defines nodes with both in-degree and out-degree equal to zero [59, 60]. *Sinks*, *sources* and *isolates* are useful model parameters for controlling degree distributions in directed networks [40]. *Out-degree (1)*, (2) and (3) are related to the *activity spread*, while *in-degree (0)*, (1) and (2) are related to *popularity spread*, as resulting from the network self-organization. *In-* and *out-degree* effects are indicative of the tendencies for centralization in the in- and out-degree distributions [41]. *Twopath* parameter accounts for the situations wherein actors who receive ties are also sending. Its model inclusion controls for the association between out-degree and in-degree.

The *relational attribute effects* in our ERGMs are predicated to affect the probability of a COVID-19 tie. The estimated effects were: ties with top three migration sending countries (*incoming migration*), ties with top three tourism sending countries (*incoming tourism*), ties with contiguous countries (*contiguity*), ties with same continent countries (*same continent*) and ties with countries wherein the language is spoken by 20% of the population (*common language 20%*).

We performed a diagnostic goodness of fit examination using the following procedure. We simulated 100 networks based on our model estimates. Then, we compared the simulated and the COVID-19 networks, examining the difference between the observed and mean scores on *dyad-wise shared partners*, *edge-wise shared partners*, *degree*, *in-degree* and *geodesic distances* [31, 40, 59–61]. In addition, we assessed whether the parameter estimates are in the ± 10 intervals and whether their standard errors do not exceed 5 [42].

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